

**On the Variance-stabilizing Non-parametric Regression  
: The case of the Nadaraya-Watson (N-W) estimator.**

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**1. Motivation.** Suppose that we wish to estimate Engel curve, which expresses the relation between income ( $X$ ) and food expenses ( $Y$ ) of a household, by one of non-parametric models, the N-W estimator. Then, we have to deal with the data that has two characteristics: One is that sample of  $X$  has distribution. Another is that the variability of  $Y$  depends on  $X$ .

**2. Introduction of the Nadaraya-Watson (N-W) estimator.** Suppose that we have  $n$ -pairs of random variables  $(X_i, Y_i)$ ,  $i = 1, \dots, n$ . We assume  $x_i$ ,  $i = 1, \dots, n$ , are the realizations of i.i.d. random variable  $X$  whose density function is denoted as  $f_X(x)$ . Let  $u_i$ ,  $i = 1, \dots, n$ , be the realizations of the disturbance random variable  $U_i$ . We assume that  $U_i|X_i$  are i.i.d. given  $X_i$  and that  $U_i$  are independent with  $X_j$ ,  $j \neq i$ . We further assume that the conditional moments of  $U|X$  as

$$E_{U|X}[U|X = x] = 0, \quad E_{U|X}[U^2|X = x] = \sigma^2(x).$$

We assume that the response  $Y$  is influenced by the explanatory variable  $X$  as

$$Y_i = m(X_i) + U_i = E(Y_i|X_i) + U_i,$$

where  $m(X_i) = E(Y_i|X_i) = \int y f_{Y|X}(y|x) dy = \int y \frac{f_{X,Y}(x,y)}{f_X(x)} dy$ . By replacing  $f_{X,Y}(x, y)$  and  $f_X(x)$  with the corresponding kernel bivariate and univariate estimates with the multiplicative kernel  $K_{X,Y}(x, y) = K_X(x)K_Y(y)$  on the numerator and the kernel  $K_X(x)$  on the denominator with the bandwidth on  $X$  as  $h_x$ , we arrive at the N-W estimator at  $X = x$  as,

$$\hat{m}_{h_x}(x) = \frac{\sum_{i=1}^n K_X\left(\frac{x-X_i}{h_x}\right) Y_i}{\sum_{i=1}^n K_X\left(\frac{x-X_i}{h_x}\right)}.$$

**3. The N-W with Mean Squared Error(MSE)-optimized bandwidth.** The bandwidth that minimizes MSE  $E_{\mathbf{X}, \mathbf{Y}}[m(x) - \hat{m}_{h_x}(x)]^2$  is obtained as,

$$h^{MSE}(x) = \left[ \frac{[\int K_X^2(t) dt] \sigma^2(x)}{[\int t^2 K_X(t) dt]^2 \frac{(2m^{(1)}(x)f_X^{(1)}(x) + m^{(2)}(x)f_X(x))^2}{f_X(x)}} \right]^{\frac{1}{5}} n^{-\frac{1}{5}}, \quad (1)$$

and the asymptotic MSE and the asymptotic variance respectively as,

$$V_{\mathbf{X}, \mathbf{Y}}(\hat{m}_{h^{MSE}(x)}(x)) \approx \left[ \frac{[\int K_X^2(t) dt]^{\frac{4}{5}} [\sigma^2(x)]^{\frac{4}{5}}}{[\int t^2 K_X(t) dt]^{-\frac{2}{5}} \frac{(2m^{(1)}(x)f_X^{(1)}(x) + m^{(2)}(x)f_X(x))^{-\frac{2}{5}}}{f_X^{-\frac{6}{5}}(x)}} \right] n^{-\frac{4}{5}}. \quad (2)$$

The N-W with MSE bandwidth has two problems. First, the variance is not uniform across the domain of the regression because the term  $\sigma^{\frac{8}{5}}(x)(2m^{(1)}(x)f_X^{(1)}(x)+m^{(2)}(x)f_X(x))^{-\frac{2}{5}}/f_X^{-\frac{6}{5}}(x)$  that appears in (2) depends on  $x$ . Second, the N-W might jump at the points where  $2m^{(1)}(x)f_X^{(1)}(x)+m^{(2)}(x)f_X(x)=0$  is satisfied because MSE bandwidth  $h^{MSE}(x)$  goes to infinity and the N-W goes to sample average  $\bar{Y} = (1/n) \sum_{i=1}^n Y_i$  at the  $x$ .

**4. The N-W with Variance-Stabilizing (V-S) bandwidth.** To solve the two problems with the MSE-optimized bandwidth, we propose the following V-S bandwidth,

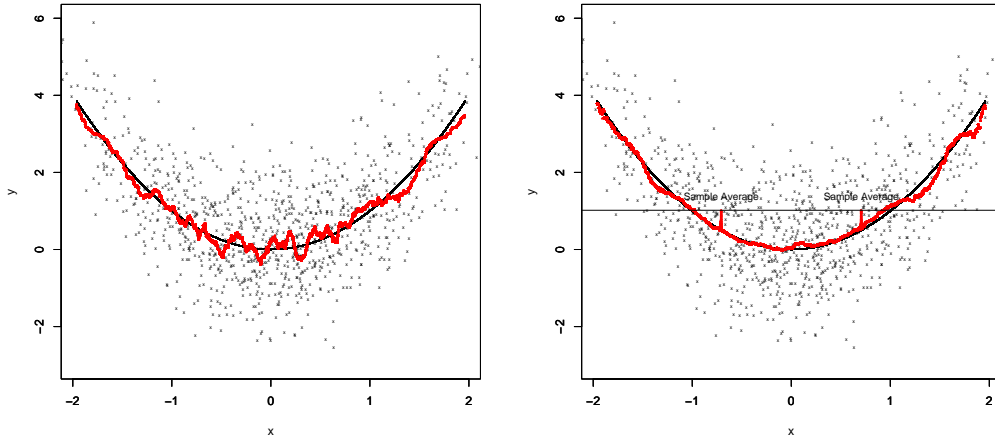
$$h^{V-S}(x) = \left[ \frac{[\int K_X^2(t)dt]}{[\int t^2 K_X(t)dt]^2 \left[ \int \frac{\sigma^8(x)(m^{(2)}(x)f_X(x)+2m^{(1)}(x)f_X^{(1)}(x))^2}{f_X^5(x)} dx \right]} \right]^{\frac{1}{5}} n^{-\frac{1}{5}} \cdot \frac{\sigma^2(x)}{f_X(x)}. \quad (3)$$

and its stabilized variance is written as,

$$V_{\mathbf{X}, \mathbf{Y}}[\hat{m}_{h^{V-S}(x)}(x)] = \left[ \frac{[\int K_X^2(t)dt]^{\frac{4}{5}}}{[\int t^2 K_X(t)dt]^{-\frac{2}{5}} \left[ \int \frac{\sigma^8(x)(m^{(2)}(x)f_X(x)+2m^{(1)}(x)f_X^{(1)}(x))^2}{f_X^5(x)} dx \right]^{-\frac{1}{5}}} \right] n^{-\frac{4}{5}} + O\left(\frac{1}{n}\right) + o\left(\frac{1}{nh_x}\right).$$

The bandwidth (3) is obtained by extracting the term  $\sigma^2(x)/f_X(x)$  from (1) and integrating the variable term except  $\sigma^2(x)/f_X(x)$  with its weight being  $f_X(x)$ . Another interpretation of (3) is the minimizer of Mean Integrated Squared Error (MISE)  $\int E_{\mathbf{X}, \mathbf{Y}}[m(x) - \hat{m}_{h_x}(x)]^2 dF_X(x)$  with its bandwidth being constrained by  $h_x = (\sigma^2(x)/f_X(x))h_0$ .

**5. Comparison of V-S with MSE bandwidth.** We show a numerical example of the case where  $m(x) = x^2 I_{[-1.96, 1.96]}(x)$ ,  $U_i|X_i \sim N(0, \sigma^2(x))$ ,  $\sigma^2(x) = 1$ ,  $X_i \sim N(0, 1)$  and  $n = 1000$ . The left panel below is the N-W with V-S bandwidth, the right with MSE-optimized bandwidth. Theoretically, variance stabilization is achieved.



The N-W with V-S bandwidth is less smooth than that with MSE-optimized. In terms of both theoretical and empirical MISE, the V-S bandwidth is inferior to MSE-optimized by about 200%. These are considered to be the penalty of variance stabilization. As for maximum error, the V-S bandwidth is superior to MSE-optimized.